Modelling Dependence in Extremal Windstorm Footprints



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The maximum 3-second wind-gust speed at each location in the 72 hour period covering the passage of the storm, centred on the time at which the maximum wind speed over land occurs



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Footprint for historical windstorm Daria (24th - 26th January 1990) wind-qust (ms^{-1}) 35 60 30 25 20 20 - 15 40 - 10 30 -20 0 20 40

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Data (kindly provided by Julia Roberts at the Met Office): footprints from the 5730 windstorms identified by Kevin Hodges' (Hodges, 1995) tracking algorithm, as occurring in the model domain in extended winters (October-March) 1979-2012. Statistical model is based on locations over land only.

Motivation

Windstorms causing extreme insurance loss have differing footprint characteristics, but which are most important for determining insurance loss?



Develop a statistical model of the windstorm footprint, to explore the relationship between these characteristics and insurance loss using sensitivity analysis techniques

Bivariate Model

By Sklar's Theorem the joint distribution of footprint wind-gust speeds at two locations, denoted X_1 and X_2 , can be though of in terms of...



Where C is often referred to as the copula function

Let
$$U = F_1(X_1)$$
 $V = F_2(X_2)$
Then, $C(u,v) = Pr(U \le u, V \le v)$

The copula contains all of the information about the joint distribution of X_1 and X_2 apart from the marginal distributions F_1 and F_2

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Independent
 $C(u, v) = uv$
 $f(u, v) = Pr(U = u, V = v) = \frac{\partial^2 C}{\partial u \partial v} = 1$
so (U,V) is
bivariate uniform
 $\int_{0}^{0} \int_{0}^{0} \int_{0}^{0$

0.8

>^{4.0}

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- Different copulas assume different extremal dependence properties
- Explore these properties in the data to select the most appropriate copula
- Extremal dependence coefficient: $\chi(u) = Pr(V > u | U > u)$

$$= Pr(V > u, U > u)/Pr(U > u)$$
$$= a/(a+b)$$



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• How does the empirical $\chi(u)$ behave for footprint wind-gusts at pairs of locations?



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- The $\chi(u)$ curves resemble the Gaussian copula
- Less dependence between London-Berlin compared to Amsterdam and Paris
- Can't calculate the limit of $\chi(u)$, $u \rightarrow 1$ empirically due to the rarity of extreme events in the data set

- Ferro (2007) proposes a probability model for the joint distribution of two variables for application in forecast verification of rare, extreme weather events (full reference on final slide)
- The model represents the diagonal of the copula as follows:

$$\chi(u) = Pr(V > u | U > u)$$
$$= \kappa (1 - u)^{\frac{1}{\eta} - 1}$$

 κ and η estimated from the data

• $0 < \eta \le 1$ is known as the coefficient of tail dependence

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 $\eta=1 \lim_{u o 1} \chi(u) > 0$ - Asymptotic dependence $\eta<1 \lim_{u o 1} \chi(u)=0$ - Asymptotic independence

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 ${f \eta}=1 \lim_{u o 1} \chi(u)>0$ - Asymptotic dependence ${f \eta}<1 \lim_{u o 1} \chi(u)=0$ - Asymptotic independence

London-Amsterdam London-Berlin London-Paris $\chi(u)$ X(U, 0.2 0.2 1.0 0.0 0.2 0.6 0.8 1.0 00 0.2 0.4 0.6 0.8 0.4 0.0 0.2 0.4 0.6 0.8 1.0 u U u

Estimate η for all pairs of locations: η < 1 (p value << 0.01)
 Footprint wind-gusts are asymptotically independent – use the Gaussian Copula

My model

- Model margins using the Generalized Extreme Value (GEV) distribution
- Model the dependence using the Gaussian Copula

Bonazzi et al. (2012) (full reference on final slide)

- Model margins using the Generalized Pareto Distribution (GPD)
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Can the models realistically represent joint losses?

Conceptual loss function: $L(X_1, X_2) = H(X_1 > t) + H(X_2 > t)$

H(a) = 1 if a is true and 0 otherwise







- Bonazzi's model, which assumes asymptotic dependence, over estimates the probability of joint extreme events
- My model does a better job of realistically representing joint losses

- Using geostatistics to develop a spatial model for all locations at once
- The spatial correlation is modeled as a function of separation distance and direction
- Quicker to simulate from compared to a multivariate copula
- Fewer model parameters, simplifying the future sensitivity analysis experiment

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 The correlation structure is similar for these locations suggesting that a geostatistical model may be able to capture the dependence structure of footprint wind-gusts.

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- Now developing a geostatistical model for windstorm footprints which models spatial dependence as a function of separation distance and direction

• Next: Finish fitting and validating the geostatistical model and carry out sensitivity analysis experiment

Thank you for listening Any questions?



References

Bonazzi, A., Cusack, S., Mitas, C., and Jewson, S. (2012). The spatial structure of European wind storms as characterized by bivariate extreme-value copulas. Nat. Hazards Earth Syst. Sci., 12:1769–1782.

Ferro, C. A. T. (2007). A Probability Model for Verifying Deterministic Forecasts of Extreme Events. Weather Forecasting, 22:1089–1100.

Hodges, K. I. (1995). Feature tracking on the unit sphere. Monthly Weather Review, 123:3458–3465.

Roberts, J. F., Champion, A., Dawkins, L., Hodges, K. I., Shaffrey, L., Stephenson, D. B., Stringer, M., Thornton, H. and Youngman, B., (2014). The XWS open access catalogue of extreme windstorms in Europe from 1979–2012. Natural Hazards and Earth System Science.



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Empirical My model Bonazzi

SSIs

Name	Date of maximum 925hPa
	windspeed over land
Daria	25 Jan 1990
Lothar	26 Dec 1999
Kyrill	18 Jan 2007
Great Storm of '87	16 Oct 1987
Vivian	26 Feb 1990
Klaus	24 Jan 2009
Martin	27 Dec 1999
Xynthia	27 Feb 2010
Anatol	3 Dec 1999
Erwin	8 Jan 2005
Herta	3 Feb 1990
Emma	29 Feb 2008
Wiebke	28 Feb 1990
Gero	11 Jan 2005
Ulli	3 Jan 2012
Dagmar-Patrick	26 Dec 2011
Fanny	4 Jan 1998
Jeanette	27 Oct 2002
Lore	28 Jan 1994
Oratio	30 Oct 2000
Stephen	26 Dec 1998
Xylia	28 Oct 1998
Yuma	24 Dec 1997





SSIs



• One such correlation function is the Matérn:

$$\rho(d) = \frac{1}{2^{\nu-1}\Gamma(\nu)} \left(\frac{2\sqrt{\nu}t}{\phi}\right)^{\nu} K_{\nu} \left(\frac{2\sqrt{\nu}t}{\phi}\right)$$

- **Γ** gamma function
- K modified Bessel function
- ϕ spatial scale parameter
- ν shape parameter added flexibility of the Matérn model
- Estimate scale parameter, ϕ , for a fixed shape, ν
- Plot over empirical binned correlogram
- Best fit for $\nu = 0.5$
- Windstorm footprints are \bigcirc_{3}^{5} a rough spatial process with correlation dropping off quickly for small separation distance







